

# Measuring Sample Quality with Stein's Method

An attempt at quantifying Monte-Carlo efficiency by Gorham and Mackey (2015)

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Our goal is to calculate quantities of the form  $\mathbb{E}_P[h(X)]$  where  $P$  is an unknown target distribution from the samples of  $Q$  as

$$\mathbb{E}_Q[h(X)] = \sum_{i=1}^n q(x_i)h(x_i).$$

We will be following (Gorham and Mackey, 2015).

- Revisiting Stein's Lemma
- Finding distances between distributions
- How Stein's method provides an answer
- Constructing Stein operators
- Calculating the discrepancies
- Experiments

# Stein's Lemma

For  $X$  mean-zero random variable and  $g$  a weakly differentiable and  $L_1$  integrable function, we have that

$$X \text{ standard Gaussian} \quad \Longleftrightarrow \quad \text{Cov}(X, g(X)) = \mathbb{E}[g'(X)],$$

$$\mathbb{E}[g'(X) - Xg(X)] = 0, \quad \longrightarrow \quad \mathbb{E}[\mathcal{A}g(X)] = 0$$

$$X \sim \text{Gamma}(\alpha, \beta) \quad \Longleftrightarrow \quad \mathbb{E}[Xf'(X) + (\alpha - \beta X)f(X)] = 0$$

$$X \sim \text{Poisson}(\lambda) \quad \Longleftrightarrow \quad \mathbb{E}[\lambda f(X+1) - Xf(X)] = 0$$

$$X \sim \text{Binomial}(n, p) \quad \Longleftrightarrow \quad \mathbb{E}[(1-p)Xf(X) - p(n-X)f(X+1)] = 0$$

$$X \sim P \quad \Longleftrightarrow \quad \mathbb{E}_P[(\mathcal{A}f)(X)] = 0.$$

# Quality Measures for samples

Target distribution  $P$  with open convex support  $\mathcal{X} \subseteq \mathbb{R}^d$ . We approximate  $P$  with  $Q$ .

Goal: to have a measure of the quality of the samples  $Q$ :

- Detects convergence  $Q_m \rightarrow P$
- Detects  $Q$  not converging to  $P$
- Is computationally feasible

A possible way to ensure these is by using an *integral probability metric*

$$d_{\mathcal{H}}(Q, P) = \sup_{h \in \mathcal{H}} |\mathbb{E}_Q[h(X)] - \mathbb{E}_P[h(X)]|$$

How to ensure that calculating this expression is tractable?

Characterizing convergence in distribution (Stein, 1972):

1. Find a real-valued operator  $\mathcal{T} : \mathcal{G} \rightarrow \mathbb{R}$  characterizing  $P$  in the sense that

$$\mathbb{E}_P[(\mathcal{T}g)(Z)] = 0 \quad \forall g \in \mathcal{G}.$$

Together,  $\mathcal{T}, \mathcal{G}$  define the *Stein discrepancy*

$$\mathcal{S}(Q, \mathcal{T}, \mathcal{G}) := \sup_{g \in \mathcal{G}} |\mathbb{E}_Q[(\mathcal{T}g)(X)]| = d_{\mathcal{G}}(Q, P).$$

2. Lower bound  $\mathcal{S}$  by some familiar IPM  $d_{\mathcal{H}}$ . **Reliability:** for  $\{\mu_m\}_{m \geq 1}$

$$d_{\mathcal{H}}(\mu_m, P) \rightarrow 0 \quad \implies \quad \mathcal{S}(\mu_m, \mathcal{T}, \mathcal{G}) \rightarrow 0.$$

3. Upper bound  $\mathcal{S}(Q, \mathcal{T}, \mathcal{G})$  to demonstrate convergence to zero (**Consistency**).

# How to construct $\mathcal{T}$

Construct a generator for a Markov process  $(Z_t)_{t \geq 0} \rightarrow P$  (Barbour, 1988). Consider a semigroup of operators  $(\mathcal{A}_t f)(x) = \mathbb{E}[f(X_t) \mid X_0 = x]$ .  $P$  is a limiting distribution if

$$\int (\mathcal{A}_t f) d\mu = \int f d\mu \quad \forall t \geq 0,$$

$$\lim_{t \downarrow 0} \int \frac{\mathcal{A}_t f(x) - f(x)}{t} d\mu(x) = 0, ,$$

$$(\mathcal{A}u)(x) = \lim_{t \rightarrow 0} \frac{1}{t} (\mathbb{E}[u(x) \mid Z_0 = x] - u(x)).$$

With this we can take for a diffusion  $dZ_t = \frac{1}{2} \nabla \log p(Z_t) dt + dW_t$  the *Stein operator*

$$(\mathcal{T}_P g)(x) \triangleq \langle g(x), \nabla \log p(x) \rangle + \nabla g(x)$$

$$\mathcal{G}_{\|\cdot\|} = \left\{ g : \mathcal{X} \rightarrow \mathbb{R}^d \mid \sup_{x,y \in \mathcal{X}} \left( \|g\|^*, \|\nabla g\|^*, \frac{\|\nabla g(x) - \nabla g(y)\|^*}{\|x - y\|} \right) \leq 1, \right. \\ \left. \langle g(x), \hat{n}(x) \rangle = 0, \forall x \in \partial\mathcal{X} \right\},$$

where

$$\|w\|^* = \sup_{\|v\|=1} \langle w, v \rangle.$$

## Observation

*This imposes conditions for all pairs of points in  $\mathcal{X}$ .*

# Bounding the Stein Discrepancy

## Lower bound:

### Theorem (*Theorem 2*)

If  $\mathcal{X} = \mathbb{R}^d$  and  $\log p$  is strongly concave with continuous and bounded 3rd and 4th derivatives then for any measures  $(\mu_m)_{m \geq 1}$ ,  $\mathcal{S}(\mu_m, \mathcal{T}_P, \mathcal{G}_{\|\cdot\|}) \rightarrow 0$  only if  $d_{\mathcal{W}}(\mu_m, P) \rightarrow 0$ .

- Sufficient but not necessary conditions for convergence.

## Upper bound:

### Theorem (*Proposition 3*)

If  $X \sim Q$  and  $Z \sim P$  with  $\nabla \log p(Z)$  integrable, then

$$\begin{aligned} \mathcal{S}(Q, \mathcal{T}_P, \mathcal{G}_{\|\cdot\|}) &\leq \mathbb{E}\|X - Z\| + \mathbb{E}\|\nabla \log p(X) - \nabla \log p(Z)\| \\ &\quad + \mathbb{E}\|\nabla \log p(Z)(X - Z)^\top\|. \end{aligned}$$

Implies convergence of  $\mathcal{S} \rightarrow 0$  whenever  $X_m \sim Q_m \xrightarrow{L^2} Z \sim P$  and  $\nabla \log p(X_m) \xrightarrow{L^1} \nabla \log p(Z)$ .



# Computing Stein Discrepancies

For observed sample values  $\{x_i\}_{i=1}^n$ , we want to solve the optimization problem

$$\mathcal{S}(Q, \mathcal{T}_P, \mathcal{G}_{\|\cdot\|}) = \sup_{g \in \mathcal{G}_{\|\cdot\|}} \sum_{i=1}^n q(x_i) (\langle g(x_i), \nabla \log g(x_i) \rangle + \nabla \cdot g_{x_i}),$$

such that  $g$  satisfies the conditions to be inside  $\mathcal{G}_{\|\cdot\|}$ .

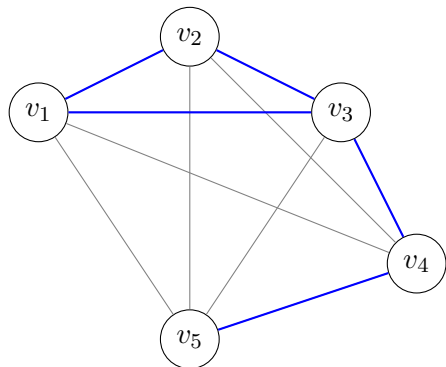
To make it feasible, Gorham and Mackey (2015) propose constraining the problem only on the values of  $g$  on  $\{x_i\}_{i=1}^n$ .

$$\sup_{\gamma_j \in \mathbb{R}^n, \Gamma_j \in \mathbb{R}^{d \times n}} \sum_{i=1}^n q(x_i) (\langle \gamma_{ji}, \nabla \log \gamma_{ji} \rangle + \Gamma_{jji}),$$

for  $\gamma_{ji} = g_j(x_i)$ ,  $\Gamma_{jki} = \nabla_k g_j(x_i)$ . An efficient way to define the constraints involves using *graph  $t$ -spanners* and an  $\ell_1$  norm.

# Graph Spanners

Graph  $G = K_n$  (Original)



—  $t$ -Spanner Edges ( $G'$ )  
— Original Edges ( $G$ )

Graph  $G'$  ( $t$ -Spanner for  $t \geq 1$ )

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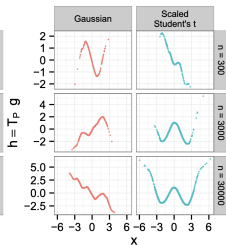
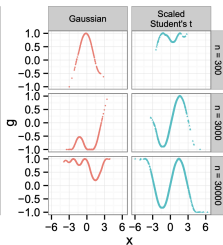
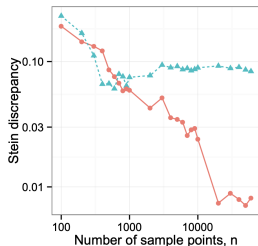
**Algorithm** Multivariate Spanner Stein Discrepancy (Algorithm 1 in Gorham & Mackey 2015)

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- 1: **input:**  $Q$ , coordinate bounds  $(\alpha_1, \beta_1), \dots, (\alpha_d, \beta_d)$
  - 2:  $G_2 \leftarrow$  Compute sparse 2-spanner of  $\text{supp}(Q)$
  - 3: **for**  $j = 1$  to  $d$  **do** (**parallelizable**)
  - 4:    $r_j \leftarrow$  Solve the  $j$ -th coordinate from linear program  $(\star)$
  - 5: **end for**  $\sum_{j=1}^d r_j$
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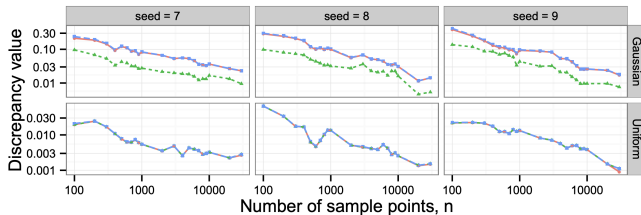
# Experiments

Target distribution  $P = \mathcal{N}(0, 1)$ .



**Sample**

- Gaussian
- Scaled Student's t



**Discrepancy**

- Classical Stein
- - - • - - - Wasserstein
- - - • - - - Complete graph Stein

- Change the diffusion process to generate  $\mathcal{T}_P$ .
- Replacing the calculation of  $\mathcal{S}$  with a Kernel Approach (Gorham and Mackey, 2017).
- Consider general diffusion operators (Gorham et al., 2019).
- If  $\mathcal{T}_P$  is too expensive to calculate, use *stochastic Stein discrepancies* (SSDs) (Gorham et al., 2020).
- And many others. . . (Anastasiou et al., 2023)

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